

Let's begin with the integral

$$\begin{aligned}
 I_n &= \frac{1}{2\pi} \int_{-\pi}^{\pi} \cos^{2n} \theta \, d\theta \\
 &= \frac{1}{2\pi} \int_{-\pi}^{\pi} \left(\frac{e^{i\theta} + e^{-i\theta}}{2} \right)^{2n} d\theta \\
 &= \frac{1}{2\pi} \int_{-\pi}^{\pi} z^{-2n} \sum_{k=0}^{2n} \binom{2n}{k} e^{ik\theta} e^{-i(2n-k)\theta} d\theta \\
 &= z^{-2n} \binom{2n}{n}. \tag{B1}
 \end{aligned}$$

Now

$$\begin{aligned}
 \sqrt{n} I_n &= \frac{\sqrt{n}}{\pi} \int_{-\pi/2}^{\pi/2} \cos^{2n} \theta \, d\theta \\
 &= \frac{1}{\pi} \int_{-\frac{\pi}{2}\sqrt{n}}^{\frac{\pi}{2}\sqrt{n}} \cos^{2n} (t/\sqrt{n}) \, dt \\
 &\xrightarrow{\text{(DCT)*}} \frac{1}{\pi} \int_{-\infty}^{\infty} e^{-t^2} \, dt = \frac{1}{\sqrt{\pi}}. \tag{B2}
 \end{aligned}$$

Comparing (B1) and (B2), we obtain

$$\lim_{n \rightarrow \infty} \sqrt{n} \binom{2n}{n} 4^{-n} = \frac{1}{\sqrt{\pi}}, \tag{B3}$$

$$\text{or} \quad \binom{2n}{n} z^{2n} \sim \frac{1}{\sqrt{\pi n}}. \tag{B4}$$

* See the end of the note for a detailed discussion.

We can also check our calculation using Stirling's asymptotic formula for $n!$:

$$\lim_{n \rightarrow \infty} \frac{n!}{\sqrt{2\pi n} n^n e^{-n}} = 1$$

or

$$n! \sim \sqrt{2\pi n} n^n e^{-n}.$$

Then

$$\begin{aligned} \binom{2n}{n} &\sim \frac{\sqrt{2\pi \cdot 2n} (2n)^{2n} e^{-2n}}{(\sqrt{2\pi n} n^n e^{-n})^2} \\ &= \frac{\sqrt{4\pi n} 2^{2n} \cancel{n^{2n}} \cancel{e^{-2n}}}{2\pi n \cdot \cancel{n^{2n}} \cancel{e^{-2n}}} \\ &= \frac{\sqrt{2}}{\sqrt{2\pi n}} = \frac{1}{\sqrt{\pi n}}. \quad (B5) \end{aligned}$$

The Central Limit Theorem II

De Moivre & Laplace noticed that the central binomial coefficients $\left\{ \binom{2n}{n+k} : -L \leq k \leq L \right\}$, where n is very big and L is small compared with n , seemed to fit a bell-shaped curve. Their analysis was similar to my earlier description of the birthday problem.

Their key insight was to consider the ratio

$$\begin{aligned}
 p_{mk} &= \frac{\binom{2m}{m+k}}{\binom{2m}{m}} = \frac{(2m)!}{(m-k)! (m+k)!} \cdot \frac{(m!)^2}{(2m)!} \\
 &= \frac{(m!)^2}{(m-k)! (m+k)!} \\
 &= \frac{m! (m(m-1) \dots (m-k+1)) (m-k)!}{(m-k)! (m+k)(m+k-1) \dots (m+1)m!} \\
 &= \frac{m(m-1) \dots (m-k+1)}{(m+k)(m+k-1) \dots (m+1)} \\
 &= \frac{\left(1 - \frac{1}{m}\right) \left(1 - \frac{2}{m}\right) \dots \left(1 - \frac{k-1}{m}\right)}{\left(1 + \frac{1}{m}\right) \left(1 + \frac{2}{m}\right) \dots \left(1 + \frac{k}{m}\right)} \quad \left. \begin{array}{l} \text{dividing top} \\ \& \text{bottom by } m \end{array} \right\}
 \end{aligned}$$

This should remind you of the birthday problem. Taking

logarithms,

$$\begin{aligned}
 \ln p_{mk} &= \sum_{j=1}^{k-1} \ln \left(1 - \frac{j}{m}\right) - \sum_{j=1}^k \ln \left(1 + \frac{j}{m}\right) \\
 &\approx -\sum_{j=1}^{k-1} \frac{j}{m} - \sum_{j=1}^k \frac{j}{m} = \frac{-k(k-1) - (k+1)k}{2m} = \frac{-k^2}{m}
 \end{aligned}$$

Hence

$$\ln p_{mk} \approx e^{-k^2/m}$$

and, if $k = t\sqrt{m}$, then $\ln p_{m, t\sqrt{m}} \approx e^{-t^2}$.

The Central Limit Theorem III

We use

$$\binom{2n}{n+zk} z^{-n} = \frac{1}{2\pi} \int_{-\pi}^{\pi} \cos^{2n} \theta e^{-z i k \theta} d\theta$$

$$= \frac{1}{\pi} \int_{-\pi/2}^{\pi/2} \cos^{2n} \theta e^{-z i k \theta} d\theta$$

$$\left(\theta = \frac{t}{\sqrt{n}} \right) = \frac{1}{\pi} \int_{-\frac{\pi}{2}\sqrt{n}}^{\frac{\pi}{2}\sqrt{n}} \cos^{2n} \left(\frac{t}{\sqrt{n}} \right) e^{-z i k t / \sqrt{n}} \frac{dt}{\sqrt{n}}$$

(B6)

Thus

$$\sqrt{n} \binom{2n}{n+2k} z^{-2n} = \frac{1}{\pi} \int_{-\frac{\pi}{2\sqrt{n}}}^{\frac{\pi}{2\sqrt{n}}} \cos^{2n} \left(\frac{t}{\sqrt{n}} \right) e^{-2ikt/\sqrt{n}} dt \quad (B7)'$$

or

$$\frac{\sqrt{n} \Gamma(2n+1) z^{-2n}}{\Gamma(n+2k+1) \Gamma(n-2k+1)} = \frac{1}{\pi} \int_{-\frac{\pi}{2\sqrt{n}}}^{\frac{\pi}{2\sqrt{n}}} \cos^{2n} \left(\frac{t}{\sqrt{n}} \right) e^{-2ikt/\sqrt{n}} dt \quad (B7)'$$

which emphasizes that both sides analytically continue.

Now, writing $2k = x\sqrt{n}$, x fixed, we obtain

$$\frac{\sqrt{n} \Gamma(2n+1) e^{-n}}{\Gamma(n+x\sqrt{n}) \Gamma(n-x\sqrt{n}+1)} = \frac{1}{\pi} \int_{-\frac{\pi}{2\sqrt{n}}}^{\frac{\pi}{2\sqrt{n}}} \cos^{2n} \left(\frac{t}{\sqrt{n}} \right) e^{-ixt} dt$$

$$\xrightarrow{\text{DCT}} \frac{1}{\pi} \int_{\mathbb{R}} e^{-t^2/2} e^{-ixt} dt = \frac{e^{-x^2/2} \sqrt{2\pi}}{\pi}$$

$$= \sqrt{\frac{2}{\pi}} e^{-x^2/2} \quad (B8)$$

or

$$\lim_{n \rightarrow \infty} \sqrt{n} \binom{2n}{n+x\sqrt{n}} z^{-2n} = \sqrt{\frac{2}{\pi}} e^{-x^2/2} \quad (B9)$$

This is a special case of the central limit theorem.

Bernoulli Trials

We say that X is a BERNOULLI RANDOM VARIABLE

if X can only take two values. For example,

if $P(X=1) = p$ and $P(X=0) = 1-p$.

A sequence of Bernoulli trials is then just a sequence of iid Bernoulli random variables X_1, X_2, X_3, \dots ; in other words, this is the mathematical idealization of tossing a coin infinitely many times. We let

$$S_n = X_1 + \dots + X_n$$

and

$$A_n = \frac{X_1 + \dots + X_n}{n}.$$

Then $E X_i = p \cdot 1 + (1-p) \cdot 0 = p$
and

$$E(X_i^2) = p \cdot 1 = p,$$

which implies

$$\begin{aligned} \text{Var } X_i &= E(X_i^2) - (E X_i)^2 \\ &= p - p^2 = p(1-p). \end{aligned}$$

Then

$$\mathbb{E} A_n = p.$$

Now

$$\begin{aligned}\mathbb{E} (A_n - \mathbb{E} A_n)^2 &= \mathbb{E} \left(\frac{X_1 + \dots + X_n}{n} - p \right)^2 \\ &= \mathbb{E} \left(\frac{X_1 - p + X_2 - p + \dots + X_n - p}{n} \right)^2 \\ &= \frac{1}{n^2} \left[\mathbb{E} (X_1 - p)^2 + \dots + \mathbb{E} (X_n - p)^2 \right] \\ &= \frac{1}{n^2} (\text{Var } X_1 + \dots + \text{Var } X_n) \\ &= \frac{p(1-p)}{n} \equiv \frac{\sigma^2}{n}.\end{aligned}$$

Hence $\{A_n\}$ is a sequence of random variables

for which $\mathbb{E} A_n = p$ and $\text{Var } A_n = \frac{\sigma^2}{n} \rightarrow 0$,

as $n \rightarrow \infty$.

Chebyshev's Inequality Let X be any random

variable with $EX = \mu$, $\text{Var } X = \sigma^2$. Then

$$P(|X - \mu| \geq \delta) \leq \frac{\sigma^2}{\delta^2}.$$

Proof:

$$\frac{\sigma^2}{n} = E(A_n - \mu)^2 \geq \delta^2 P(|A_n - \mu| \geq \delta) \quad \square$$

WEAK LAW OF LARGE NUMBERS: For any $\delta > 0$,

$$P(|A_n - \mu| \geq \delta) \leq \frac{\sigma^2}{n \delta^2}.$$

Proof:

$$\frac{\sigma^2}{n} = E(A_n - \mu)^2 \geq \delta^2 P(|A_n - \mu| \geq \delta)$$

\square

BERNSTEIN POLYNOMIALS

Define

$$\begin{aligned} B_n f(p) &= \mathbb{E} f(A_n) \\ &= \sum_{k=0}^n f\left(\frac{k}{n}\right) \binom{n}{k} p^k (1-p)^{n-k}. \end{aligned}$$

Then $B_n f$ is a polynomial of degree n . Since A_n is likely to be close to $f(p)$. In fact much more is true. First observe that

$$\begin{aligned} |f(p) - B_n f(p)| &= \left| \mathbb{E} \left(f(p) - f(A_n) \right) \right| \\ &\leq \mathbb{E} |f(p) - f(A_n)| \end{aligned}$$

Now suppose that $f: [0, 1] \rightarrow \mathbb{R}$ is a continuous function. A theorem in analysis states that any such f is UNIFORMLY CONTINUOUS: given any $\epsilon > 0$, there is a $\delta > 0$ for which $|x - y| \leq \delta$ implies $|f(x) - f(y)| \leq \epsilon$, for any $x, y \in [0, 1]$.

Then

$$|f(p) - B_n f(p)|$$

$$\leq 2 \|f\|_\infty \mathbb{P}(|A_n - p| \geq \delta)$$

$$+ \epsilon \mathbb{P}(|A_n - p| < \delta)$$

$$\leq \frac{2 \|f\|_\infty \sigma^2}{n \delta} + \epsilon.$$

Hence, given $\epsilon > 0$, choosing an enormous n

implies

$$|f(p) - B_n f(p)| \leq 2\epsilon$$

for any $p \in [0, 1]$.

